Some Results on the Statistics of Delay Terms in SR ARQ on Markov Channels

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Abstract—In this paper we explore the packet delay statistics of a Selective Repeat ARQ scheme on Discrete Time Markov Channel with non-instantaneous round trip delay. In particular, we are interested in obtaining considerations about the queueing delay of the process and also possible comparisons between different delay components. For this reason, we analyze in detail the impact of system parameters, such as the packet arrival rate and the packet error probability, on the terms which constitute the overall delay. Finally, we explore the connection of these numerical evaluations with the QoS requirements connected to delay for multimedia traffic.

Index Terms—Automatic repeat request, Markov processes, error analysis, delay estimation.

I. INTRODUCTION

THE study we present in this paper deals with fine tuning of the system parameters in SR ARQ systems, in order to better understand the trade-off between the terms which constitute the overall delay experienced by a packet.

Selective Repeat Automatic Retransmission reQuest (SR ARQ) [1] is an error control technique, where the retransmission of negatively acknowledged packets is selectively triggered by the receiver, so that after a retransmission the data flow is resumed from the last packet sent so far. In our analysis the time for a packet transmission corresponds to one slot and feedback packets, containing either an acknowledgement (ACK) or a not-acknowledgement (NACK) messages come back at the transmitter after a full round trip time. Data packets are released *in-order* to higher layers, i.e., release is possible once every packet with lower identifier has been acknowledged. The receiver keeps in a buffer the packets correctly received but not yet released, so that the sender retransmits not acknowledged packets only.

The main aspect of our investigation, which enforce the choice of SR ARQ as representative also of other ARQ-like techniques, is that we assume that the transmission feedback, expressed by the receiver with acknowledgement / not acknowledgement (ACK/NACK) packets, is not instantaneous at transmitter's side. Between an erroneous transmission and the corresponding retransmission other packets are sent over the channel. Since data packets must be released to higher layers *in-order*, i.e., when every packet with lower identifier has been acknowledged, the receiver keeps in a buffer the packet correctly received but not yet released.

Thus, the overall delay term is subdivided into different terms: at the transmitter, a queueing delay is experienced by

the packets, which are also delayed by retransmissions, which achieve higher priority. Also after transmission, at the receiver packets are released after the correct reception of every packet with lower identifier. To the *transmission delay* of a packet an additional term must be added, which takes into account the transmission of previous packets which are still to be acknowledged. The delay spent by a packet in the receiver buffer waiting for the final delivery is referred to as *re-sequencing delay*. The sum of the transmission and re-sequencing delay is also called *delivery delay*. These subdivisions are quite standard in the literature [2] and reflect different aspects of the system.

The transmission of multimedia data, which often is performed by means of ARQ-like techniques, is generally considered to be subject to constraints related to the Quality of Service (QoS). Henceforth, a precise understanding of the delay performance is required. In fact, delay terms can be directly connected with other specific QoS issues. In particular, the queueing delay can be related with the transmission buffer occupancy, where the delivery delay has the same role for the receiver buffer [3]. Moreover, for Next Generation services it is expected to have not only throughput requirements (which is the case for data traffic) but also real time constraints, and not only on the average delay but also on the jitter [4], [5]. Thus, a better understanding of delay statistics with analytical instruments is worth of interest, since it is able to capture in an exact way parameters which are difficult to characterize otherwise. In particular, we characterize also, as clearly allowed by our model, quantities related to second order moments, which somehow relate to the delay jitters.

Another important aspect which is often neglected when performing analysis of delay performance is that not only wireless channels have a generally higher error rate than wireline ones, but there is an additional factor which heavily affects the performance, i.e., the error correlation. Hence, we are interested in considering error correlation in our study, since it has been proven that the performance of ARQ protocols can change dramatically [6], [7]. A suitable way, which permits analytical investigations, is the employ of a Markov chain to represent the underlying channel [8]. In fact, channel burstiness can easily be introduced by appropriately defining the elements of the transition probability matrix.

This is important since, as we will show quantitatively, the impact of channel errors can not be taken into account by looking at the average packet error rate only. Moreover, channel burstiness has a heavy impact on the performance, being in particular the queueing delay heavily underestimated when an *iid* error process is considered instead of a correlated one. Also, the standard deviation of the queueing delay increases even more rapidly when the channel errors are strongly correlated.

These conclusions are obtained by exploiting an analytical model which directly stems by our previous papers presented in [9], [13]. The contribution here is to extend the analysis in order to show numerical results and describe in detail the impact of system parameters. In this way, several practical conclusions can be drawn for the setup of ARQ-like systems, which we believe are difficult to estimate precisely without an exact analysis.

The rest of this paper is organized as follows: in Section II we discuss other research contributions on topics related to the delay statistics of ARQ systems. In Section III we discuss the system model, with particular focus on the key system parameters. In Section IV we present numerical results with a detailed discussion about the higher order statistics of the queueing delay and the comparison of different delay terms. Finally, Section V concludes the work.

II. RELATED WORK

Several papers can be related with the delay performance of SR ARQ. In [10], an analytical model for the packet delay and buffer occupancy in a SR ARQ system has been proposed for a static channel (fixed error probability), where [11] and [3] analyzed in detail the re-sequencing terms, again in the independent error case. The impact of channel correlation for different ARQ techniques (i.e., also including Stop-and-Wait, Go-Back-N, adaptive SR ARQ and ideal SR ARQ) was considered in [5]-[8], [12]. The separate analysis and comparison of different delay terms for a time varying channel has been considered in [2], but only quantifying approximate mean values. In [9], [13] we investigated instead Markov techniques to derive delay statistics (delivery and queueing delay, respectively) in an exact manner. In particular, both in [13] and in the present paper we adopted a Bernoulli arrival process to characterize the packet generation. This was introduced for the first time in [14]. Instead, in [4] a more general packet arrival process is considered. This last contribution and a very recent paper [15] also investigate the queueing delay in detail.

III. SYSTEM MODEL

We consider a pair of communicating entities (a transmitter and a receiver) that exchange packets through a noisy wireless link and use an SR ARQ transmission technique with unlimited re-transmission attempts. The time for a packet transmission corresponds to one slot and feedback packets, containing either an acknowledgement (ACK) or a not-acknowledgement (NACK) messages, come back at the transmitter after a full round trip time, which equates m slots. We focus on the more interesting case of non-instantaneous feedback, i.e., m > 1. Fig. 1 shows the way in which packets are stored and transmitted, and the consequent delays.

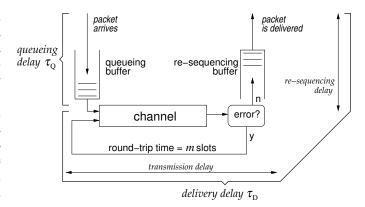


Fig. 1. The model of SR ARQ and its delays

Note that the transmitter continuously transmits new packets in increasing numerical order as long as ACKs are received. In case of NACK, the packet must be retransmitted, which happens with higher priority with respect to the packets of the queue. However, as ACKs/NACKs refer to the transmission of m slots before, the numerical increasing order of the identifier of the transmitted packets might be broken. This explains why the delay experienced by a packet is related also to the outcome of other packets, due to out-of-order transmissions and re-sequencing.

As will be exploited in the following, this implies that, since in every case the bottleneck of previous packets still pending and henceforth blocking the delivery of the ones with higher id can not exceed m slots, we need to keep track of the last m transmissions.

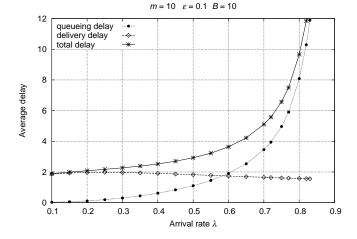
For the analysis purpose, we assume to have unlimited buffer capacity both for the queue at transmitter's side and also for the re-sequencing buffer at receiver's side. We also focus on the case of error-free ACK/NACKs. These are quite common assumptions, considered only for the sake of clarity, but they do not substantially change the analysis.

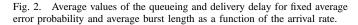
We also adopt a Markov model to represent the channel state [16]. In this way, we are able to describe the transmission with more variables than the error probability value alone. The Markov channel we adopt here is Two-State, i.e. state 0 corresponds to an error-free channel condition, where state 1 is always erroneous. The channel transition probability matrix is then:

$$\mathbf{P} = \begin{pmatrix} p_{00} & p_{01} \\ p_{10} & p_{11} \end{pmatrix}, \tag{1}$$

and the average channel error probability and the average burst length are therefore $\varepsilon=p_{01}/(p_{10}+p_{01})$ and $B=1/p_{10}$, respectively. If $B=\varepsilon^{-1}$ channel errors are i.i.d., otherwise (usually if $B>\varepsilon^{-1}$) there is error correlation. In the following we will refer to these parameters, which completely describe the channel transition matrix, due to their better clarity from the description point of view. Even this simple model accounts for the channel correlation with the average error burst length B, whereas on the other hand offers the advantage of describing correlation with only a single parameter. Of course, more complicated Markov channels can be used as well if necessary.







For the arrival process, a Bernoulli model is considered, where a packet arrival occurs with probability λ during a slot. We have chosen this mechanism since it allows a simple representation with a single parameter of different load conditions. Of course a more complicated model, e.g., with correlated arrivals also, would be more realistic, but preliminary investigations have shown that for the kind of insight we intend to give in the following, this model is more than appropriate.

Note that, for the system to admit a convergent steady state distribution, λ must be between 0 and $1-\varepsilon$, that is the channel service rate. When $\lambda>1-\varepsilon$, τ_Q becomes indefinite, since the queueing buffer is saturated. This condition, called in the literature $Heavy\ Traffic\ [2]$, would anyway allow for an analysis of the delivery delay only, as in [9]. Indeed, as we will show in the following, the Heavy Traffic condition probably suffices for most practical purposes when studying the delivery delay, since the sensitivity of the delivery delay on the arrival intensity is rather weak. However, the most interesting investigations performed in this paper, which involve the queueing delay and/or the overall delay, can be performed only if a variable packet arrival rate is accounted for.

This model can be solved for example by following a Matrix-Geometric approach [17], as we already performed in [13]. For this reason, we do not present the model resolution here, if not in the form of a brief summary. Interested readers might follow the details of the analysis, with also proof of correctness and comparison with simulation results, in the aforementioned paper. Note that the contribution here is instead to show how this model can be actually applied to derive considerations about the system dimensioning. In particular, in the following section we will use it to compare different delay terms and also to quantify the queueing delay effect, which, thanks to the analytical evaluation, can be done not only through average values but also with higher order moments.

The model resolution can be directly derived from the following observations. First of all, the system memory comprises the current queueing buffer occupancy and the current channel

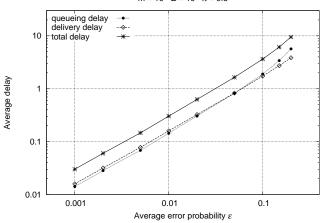


Fig. 3. Average values of the queueing and delivery delay for fixed average burst length and arrival rate as a function of the average error probability.

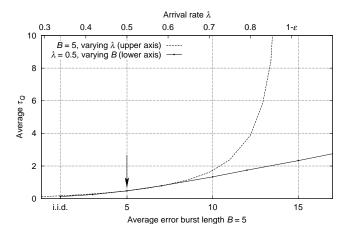
state. The information about past events can be instead fully characterized by the outcome of last m transmissions. It is in fact true that the transmissions and the deliveries might be blocked by an error occurred in the transmission arbitrarily far in the past. However, the corresponding packet has to be re-transmitted at most within m-1 slots before the packet currently on the channel.

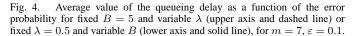
This implies that an appropriate Markov chain can be defined, where part of the state relates to the channel evolution, hence, it follows the channel DTMC. Another term in the system state derives from the queueing buffer state and evolves according to the following alternatives: if a packet arrives and a retransmission is scheduled, the queue size increases by one; if no packet arrives and no retransmission is scheduled, the buffer occupancy decreases by one, provided it is not zero already; in every other case it stays the same. All these conditions are easily mapped through an appropriate transition matrix of the Markov chain. Finally, the system state comprises also the information about the last m transmissions, which evolves deterministically: for this reason, it is simply a technical matter to put the transitions between the appropriate states. Thus, after some calculus, one can determine the steady-state values of the system distribution and hence evaluate the delay statistics, as explained in [13].

IV. RESULTS

The first results we show in this contribution focus on the comparison of delay terms. An example is shown in Fig. 2. Here, the two main parts of the overall delay are compared, as a function of the arrival rate. This figure refers to a correlated channel with average burst length comparable with the total round-trip-time, since both m and B are taken equal to 10 slots. The average error probability ε has been set to 0.1.

The figure emphasizes that the queueing delay rapidly increase with λ , so that the queueing and the delivery delays are more or less of the same order of magnitude when λ is between 0.5 and 0.7, whereas the queueing delay is as a matter of fact negligible for $\lambda < 0.5$ and explodes for values higher than 0.7.





Therefore, the arrival process mostly affect the queueing term; indeed, the delivery delay is also impacted, but in a more complicated manner that will not be discussed here in detail. A more thorough description of this behavior can be found in [18]. However, for the considered values the average delivery delay remains almost constant. This last property has been verified to hold every time the channel burstiness is comparable with m.

Another similar comparison is reported in Fig. 3, where the average delays are evaluated as a function of the average error probability ε . In this case, we choose λ equal to 0.6, whereas m and B are both still equal to 10. Also this figure shows that, as the channel conditions becomes more restrictive, the queueing delay impact becomes predominant. In fact, for low error probabilities τ_Q and τ_D are more or less comparable, and therefore the overall delay is roughly twice each of them. However, while the delivery delay increases more or less linearly with ε , the queueing delay follows a similar behavior but explodes when the channel error probability is so high that the Heavy Traffic condition is approached. In other words, apart from the cases the arrival intensity is so high that the queueing buffer is stuck, all delay terms increases more or less linearly with λ .

For the delivery delay this proportional increase keeps holding even for very high arrival rates. This confirms again that the delivery delay can be appropriately studied even under the Heavy Traffic assumption, whereas the queueing delay obviously can not. For this reason, in the following results we mainly focus on the queueing delay and its relationship with the arrival process.

Another important aspect to investigate is the channel burstiness. Whereas in [9] we already proved that the evaluation of the delivery delay statistics can be accurate only if the channel correlation is taken into account, in this paper we quantitatively support a similar statement for the queueing delay also. Even though bursty and independent channels might exhibit a similar qualitative behavior, since of course as shown before τ_Q is strongly dependent on the arrival and departure rate of the system, an iid channel is definitively not

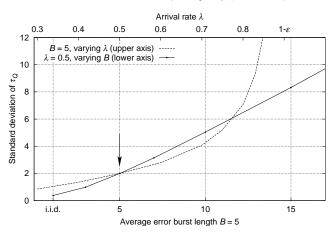


Fig. 5. Standard deviation of the queueing delay as a function of the error probability for fixed B=5 and variable ε (upper axis and dashed line) or fixed $\varepsilon=0.1$ and variable B (lower axis and solid line), for $m=7, \lambda=0.5$.

a good model for correlated wireless channels from the point of view of evaluating the delay statistics. For example, the tails of the distributions for bursty channels are heavier, which means that long burst of errors might let the queueing delay significantly increase.

This can be shown in Fig. 4, where the mean τ_Q is evaluated for different values of the average error burst length B (solid line). In this figure, m=7 and $\varepsilon=0.1$ are considered. The lower value of m is simply motivated by the faster computational evaluation. For comparison, also the dependence on the arrival rate λ is considered, with the dashed curve, which follows the same behavior of the queueing delay curve in Fig. 2. This figure should be read by keeping in mind that the middle value of both curves (indicated by the arrow) refers to the same case, and moving on the solid or dashed curve means to change B or λ , respectively. In this way, it is possible to see that the queueing delay is clearly larger in the correlated channel than in the case of i.i.d. errors. The increase of the queueing delay can be considered linear in B, whereas it explodes as λ increases and approaches $1-\varepsilon$.

Since we took an analytical approach, we are also able to investigate, e.g., second order moments of the statistics, which have a non negligible impact on the subjective perception of the QoS due to delay jitters. The standard deviation of the queueing delay is plotted in Fig. 5, where we adopt the same way of representing the data than Fig. 4 in order to focus on and compare the dependance of this value on λ and B. What is interesting to observe, is that the standard deviation of τ_Q is also increasing linear in B. This hold for all values but for the i.i.d. case, where the behavior is slightly different. This again confirms that considering the channel correlation is unavoidable if a correct evaluation of the statistics is required.

Moreover, note that for what concerns the delay jitter the impact of B is even more relevant than for the average value. In fact the increase in the standard deviation of τ_Q when B is changed from 5 to 10 is comparable with the variation which the same value would have observed if λ would be increased instead to almost 0.8, which is a significantly higher value. In other words, Fig. 5 shows that the delay jitter of a

heavily bursty channel are the same of a system with lower B but higher arrival rate. Even though this can be intuitively explained by considering that the higher B the more frequent the retransmissions, which is eventually similar to a higher arrival rate, Fig. 5 shows that nevertheless this phenomenon severely impacts on the delay. This conclusion can be directly connected to the QoS requirements in terms of queueing buffer and delay jitter for real-time traffic, indicating that the negative impact of the error correlation is times heavier than what one might expect by looking at the average value only.

V. CONCLUSIONS

In this paper we have studied the delay performance of a Selective Repeat ARQ scheme over a Markov Channel. The statistics obtained with an analytical model have been used to evaluate the single delay terms and their main characteristics are compared for several values of the channel error probability and error correlation.

The following interesting conclusions can be inferred from the numerical evaluations. First of all, the queueing delay is very sensitive on the arrival process intensity, so that is generally increasing as the packets arrive more frequently, until it reaches instability of course when the channel saturation is approached. This might seem a trivial conclusion, since it is somewhat expected that a higher packet arrival rate makes both queue and queueing delay longer. However, when the channel correlation is considered, this increase is amplified. In particular, the second-order analysis reveals that this effect is even more relevant for the delay jitters, so that avoiding heavily correlated channel errors appears as a key point in order to meet real-time traffic QoS constraints.

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